

Stochastic Modelling and Forecasting for LPG Prices: SARIMA Approach

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Abstract

Liquefied Petroleum Gas (LPG) is a by product of crude oil and natural gas that is rich in hydrocarbons. LPG can be used for cooking, water heating, and space heating and drying. LPG allows the user to control the level of heat and ensures a clean cooking environment. This study examines to predict and forecast analysis for Chennai's monthly LPG prices from December 2013 to November 2021 through the Seasonal Autoregressive Integrated Moving Average (SARIMA) approach. The use of Mean Absolute Percentage Error (MAPE) and Bayesian Information Criterion (BIC) criteria for the fitted model considered appropriate LPG pricing. The best model to forecast the Gas Prices is SARIMA(0,1,1)(1,1,0)₁₂. Based on the chosen model, it could be predicted that LPG price would increase to Rs. 1,308.75 in November 2023 from Rs. 900.20 in December 2021 in Chennai.

Key Words: LPG Price forecasting, PACF, SARIMA.

Introduction

LPG is a multipurpose fuel and used worldwide as an alternative fuel for various applications; residential, commercial, agribusiness, industrial, and auto gas. In the absence of natural gas, LPG is the best alternative fuel. Cooking with other fuels such as wood or kerosene results in pollution, accumulation of soot, and the heat generated is difficult to control. LPG has been in India since 1950 and was used only by the middle and upper class, despite government subsidies for all. In recent years, the Government of India has sought to change these access and utilization practices through a series of target policies. From 2009 to 2012, Rajiv Gandhi Gramin LPG Vitaran Yojana provided 1.5 million new LPG connections to rural areas (Jain, 2016). Since 2015, the Government of India, in collaboration with three major oil companies, has launched three major projects to improve LPG for poor and rural households. LPG prices in India are based on the Import Parity Price (IPP) of the international market and the fuel is imported into the country.

Materials and Methods

As the aim of the study was to predict and forecast analysis for Chennai's monthly LPG prices from December 2013 to November 2021 through the Seasonal Autoregressive Integrated Moving Average (SARIMA) approach, and data collected from the Annual Report (2020-21) of Indian Oil Corporation Limited. Akaike (1970) discussed the stationary time series by an AR (p), where p is finite and bounded by the same integer. Moving Average (MA) models were used by Slutzky (1973). Box and Pierce (1976) considered the distribution of residual autocorrelations in ARIMA. SARIMA model introduced by Box and Jenkins (1976), was frequently used for discovering the pattern and predicting the future values of the time series data. Hannan and Quinn (1979) suggested obtaining the order of a time series model by minimizing the errors for pure AR models, and Hannan (1980) for ARMA models. A second order determination method could be considered as a variance of Schwarz's Bayesian Criterion (SBC) which gives a consistent estimate of the order of an ARMA model. Hosking (1981) introduced a family of models, called fractionally differenced autoregressive integrated moving average models, by generalizing the 'd' fraction in ARIMA (p,d,q) model. Pindyck (1994) explored the relationship between short-term stocks and prices for copper, wood, and heating oil. Solomon (2001) sought to predict future oil and gas prices by triggering a linear regression analysis of the impact of oil and natural gas prices on oil industry activity. Stephane et al., (2003) attempted to estimate the actual oil price. Zamani (2004) proposed an economic forecasting model for short-term oil spot prices. Japlonowski et al., (2007) proposed a conclusion analysis model for evaluating the crude oil price forecast. Ogwo (2007) proposed an equivalent gas price model. Mogopadhyay et al., (2012) suggested that fuel stacking practiced for cooking methods in rural homes with

LPG. Holoda et al., (2017); Troncoso and da Silva, (2017) explained an important task of the fuel stacking procedures for the popular and clean option.

The seasonal ARIMA model incorporates both non-seasonal and seasonal factors in a multiplicative model. One shorthand notation for the model is

$$\text{ARIMA}(p,d,q) \times (P,D,Q)_s$$

with p = non-seasonal AR order, d = non-seasonal differencing, q = non-seasonal MA order, P = seasonal AR order, D = seasonal differencing, Q = seasonal MA order and S = time span of repeating seasonal pattern, Φ = Autoregressive polynomial Order and Θ = Moving average polynomial Order.

Without differencing operations, the model could be written more formally as

$$\Phi(B^S)\phi(B)(x_t - \mu) = \Theta(B^S)\theta(B)\omega_t \text{ ----- (1)}$$

The non-seasonal components are:

$$\text{AR: } \phi(B) = 1 - \phi_1 B - \dots - \phi_p B^p$$

$$\text{MA: } \theta(B) = 1 + \theta_1 B + \dots - \theta_q B^q$$

The seasonal components are:

$$\text{Seasonal AR: } \Phi(B^S) = 1 - \Phi_1 B^S - \dots - \Phi_p B^{pS}$$

$$\text{Seasonal MA: } \Theta(B^S) = 1 + \Theta_1 B^S - \dots - \Theta_Q B^{QS}$$

On the left side of equation (1) the seasonal and non-seasonal AR components multiply each other, and on the right side of equation (1) the seasonal and non-seasonal MA components multiply each other.

SARIMA Analysis:

Table 1: Actual LPG Price (Rs.) in Chennai

Month	Amount	Month	Amount	Month	Amount	Month	Amount	Month	Amount
Dec 2013	1014	Aug 2015	604	Apr 2017	732	Dec 2018	827	Aug 2020	611
Jan 2014	1234	Sep 2015	577	May 2017	639	Jan 2019	705	Sep 2020	610
Feb 2014	1132	Oct 2015	532	Jun 2017	560	Feb 2019	673	Oct 2020	610
Mar 2014	1082	Nov 2015	560	Jul 2017	574	Mar 2019	717	Nov 2020	610
Apr 2014	981	Dec 2015	623	Aug 2017	533	Apr 2019	722	Dec 2020	685
May 2014	929	Jan 2016	672	Sep 2017	609	May 2019	728	Jan 2021	710
Jun 2014	905	Feb 2016	587	Oct 2017	657	Jun 2019	753	Feb 2021	760
Jul 2014	924	Mar 2016	526	Nov 2017	750	Jul 2019	653	Mar 2021	835

Aug 2014	922	Apr 2016	521	Dec 2017	756	Aug 2019	591	Apr 2021	825
Sep 2014	903	May 2016	538	Jan 2018	751	Sep 2019	607	May 2021	825
Oct 2014	883	Jun 2016	560	Feb 2018	746	Oct 2019	620	Jun 2021	825
Nov 2014	864	Jul 2016	551	Mar 2018	700	Nov 2019	696	Jul 2021	851
Dec 2014	750	Aug 2016	500	Apr 2018	664	Dec 2019	714	Aug 2021	876
Jan 2015	705	Sep 2016	478	May 2018	663	Jan 2020	734	Sep 2021	901
Feb 2015	600	Oct 2016	501	Jun 2018	713	Feb 2020	881	Oct 2021	916
Mar 2015	606	Nov 2016	539	Jul 2018	771	Mar 2020	826	Nov 2021	916
Apr 2015	614	Dec 2016	594	Aug 2018	806	Apr 2020	762		
May 2015	609	Jan 2017	595	Sep 2018	839	May 2020	570		
Jun 2015	620	Feb 2017	661	Oct 2018	896	Jun 2020	607		
Jul 2015	628	Mar 2017	747	Nov 2018	960	Jul 2020	611		

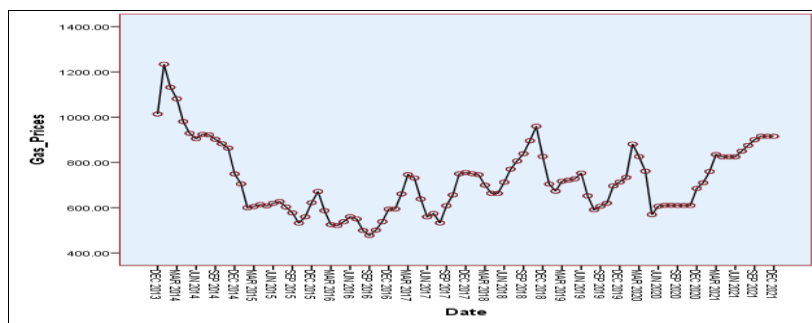


Figure1: Time Plot for Month-wise LPG Prices

Figure 1 depicts that the data used were non-stationary. The time series plot of monthly Gas Prices from December 2013 to December 2021 with $d=1$ was created as shown in Figure 2.

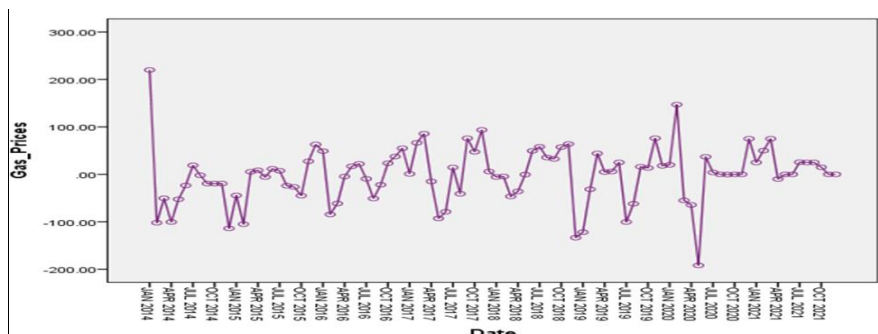


Figure 2: Time Plot for Month-wise LPG Prices with $d=1$

Table 2: ACF and PACF of LPG Prices with d=1

Lag	Autocorrelation	Std. Error ^a	Box-Ljung Statistic			Partial Autocorrelation	Std. Error
			Value	df	Sig. ^b		
1	0.221	0.100	4.843	1	0.028	0.221	0.102
2	0.007	0.100	4.848	2	0.089	-0.044	0.102
3	-0.241	0.099	10.706	3	0.013	-0.245	0.102
4	-0.142	0.099	12.772	4	0.012	-0.039	0.102
5	-0.109	0.098	13.995	5	0.016	-0.073	0.102
6	-0.073	0.098	14.559	6	0.024	-0.104	0.102
7	0.017	0.097	14.591	7	0.042	0.014	0.102
8	0.137	0.097	16.611	8	0.034	0.100	0.102
9	0.139	0.096	18.710	9	0.028	0.046	0.102
10	0.126	0.096	20.450	10	0.025	0.084	0.102
11	-0.061	0.095	20.857	11	0.035	-0.066	0.102
12	-0.031	0.094	20.966	12	0.051	0.050	0.102
13	-0.162	0.094	23.953	13	0.032	-0.106	0.102
14	-0.087	0.093	24.813	14	0.036	-0.033	0.102
15	0.105	0.093	26.090	15	0.037	0.176	0.102
16	0.123	0.092	27.871	16	0.033	0.015	0.102

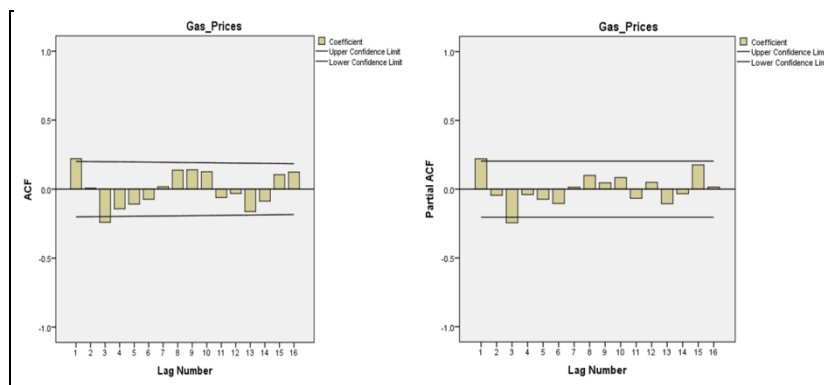


Figure 3: ACF and PACF plot with d=1

There is a significant autocorrelation at 3 lags which indicates that seasonality of data needed to be considered (Figure 3). First seasonal differencing needed to be done. The time series plot of monthly Gas Prices with $d=1$ and $D=1$ was created as shown in Figure 4. Stationary and normal time series plot with seasonality proof after the transformation of difference, $d=1$ and seasonal difference, $D=1$. The ACF and PACF plots with transformation $d=1$ and $D=1$ were depicted as shown in Figure 5.

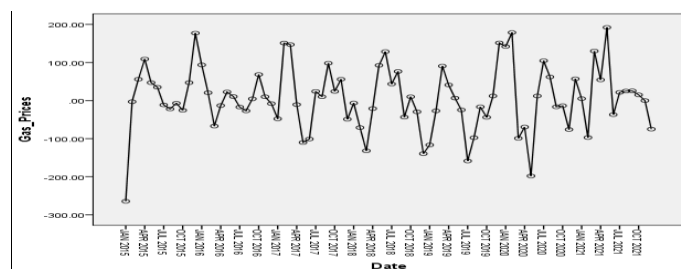


Figure 4: Time Plot for Month-wise LPG Prices with d=1 and D=1

Table 3: ACF and PACF of LPG Prices with $d=1$ and $D=1$

Lag	Auto correlation	Std. Error ^a	Box-Ljung Statistic			Partial Auto correlation	Std. Error
			Value	df	Sig. ^b		
1	0.312	0.107	8.493	1	0.004	0.312	0.109
2	-0.051	0.107	8.725	2	0.013	-0.165	0.109
3	-0.401	0.106	23.102	3	0.000	-0.377	0.109
4	-0.287	0.105	30.548	4	0.000	-0.066	0.109
5	-0.135	0.105	32.221	5	0.000	-0.087	0.109
6	-0.047	0.104	32.423	6	0.000	-0.197	0.109
7	0.029	0.103	32.500	7	0.000	-0.076	0.109
8	0.171	0.103	35.293	8	0.000	0.108	0.109
9	0.300	0.102	43.977	9	0.000	0.176	0.109
10	0.202	0.101	47.954	10	0.000	0.050	0.109
11	-0.146	0.101	50.068	11	0.000	-0.179	0.109
12	-0.443	0.100	69.764	12	0.000	-0.260	0.109
13	-0.214	0.099	74.425	13	0.000	0.146	0.109
14	-0.031	0.098	74.521	14	0.000	-0.067	0.109
15	0.216	0.098	79.412	15	0.000	-0.015	0.109
16	0.151	0.097	81.825	16	0.000	-0.005	0.109

^aThe underlying process assumed is independence (white noise).

^bBased on the asymptotic chi-square approximation.

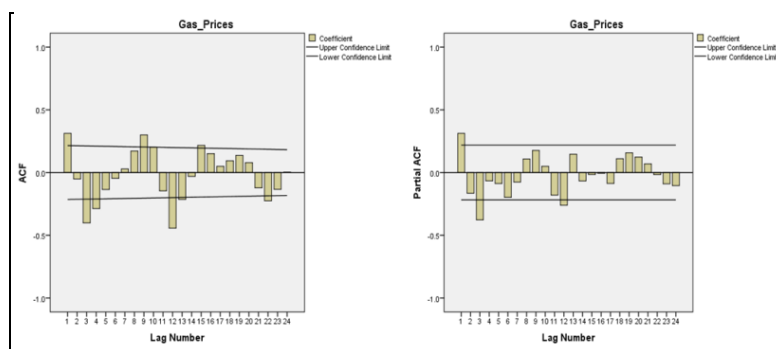


Figure 5: ACF and PACF plot with $d=1$ and $D=1$

From the ACF plot (Figure 5), there is significant autocorrelation at Lags 1 and 3 which will be the value of q . While the significant autocorrelation at Lag 3 indicates that the value of $Q=1$. From the PACF plot, there is significant autocorrelation at Lags 1 and 3 which will be the value of p . While the autocorrelation at Lag 3 presents but less significant, hence, it indicates that the value of $P = 0$ or $P = 1$. In addition, the PACF trails off after a lag and has a hard cut-off in the ACF after the lag q , at $p=0$ needed to be considered in the models.

In this study, the lags 12 on the order AR and MA were ignored. Table 4 displays parameter estimator (column 3) and the significance test of α value (column 4). The following hypothesis will be used to identify significant models:

$$H_0: \alpha > 0.05: \text{Model insignificant}$$

$$H_1: \alpha < 0.05: \text{Model significant}$$

There were 6 initial models being identified from Table 4, it is clearly shown that only SARIMA (0,1,1)(1,1,0)₁₂ has reached the assumptions of the model diagnostics.

Table 4: Significant parameters of identified models

Model: SARIMA	Parameters	Estimation	Significant	MAPE	Normalized BIC
(1,1,0)(1,1,0) ₁₂	AR	0.373	0.001	7.703	8.692
	SAR	-0.564	0.000		
(1,1,1)(1,1,1) ₁₂	AR	0.216	0.552	7.318	8.749
	MA	-0.100	0.784		
	SAR	-0.224	0.274		
	SMA	0.525	0.009		
(1,1,1)(1,1,0) ₁₂	AR	0.303	0.290	7.277	8.674
	MA	-0.095	0.745		
	SAR	-0.564	0.000		
(0,1,1)(1,1,0)₁₂	MA	-0.318	0.003	7.352	8.625
	SAR	-0.554	0.000		
(0,1,1)(1,1,1) ₁₂	MA	-0.266	0.014	6.885	8.584
	SAR	-0.181	0.349		
	SMA	0.613	0.002		
(1,1,0)(1,1,1) ₁₂	AR	0.301	0.006	7.349	8.687
	SAR	-0.228	0.250		
	SMA	0.526	0.008		

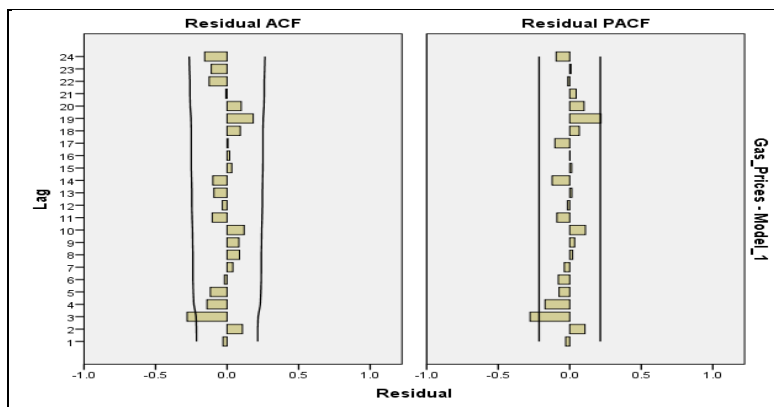


Figure 6: Residuals of ACF and PACF

The residual plots of ACF and PACF in Figure 6 shows that, the model is adequate a random variation from the origin zero (0), the points below and above are all uneven, hence the model fitted is adequate.

Model parameters and fit statistics were estimated and the results of estimation are presented in Tables 5. BIC value was 8.625. Hence, the most suitable model for LPG prices was SARIMA (0, 1, 1) (1, 1, 0)₁₂, as this model had the lowest normalized BIC value.

Table 5: ARIMA Model Parameters

		Estimate	SE	t	Sig.
Constant		4.010	6.636	0.604	0.547
Difference		1			
MA	Lag 1	-0.312	0.104	-3.002	0.004
AR, Seasonal	Lag 1	-0.551	0.100	-5.510	0.000
Seasonal Difference		1			

The forecasted plot of SARIMA(0,1,1)(1,1,0)₁₂ through Box-Jenkins method is shown in Figure 7 and the forecasted results were tabulated in Table 6.

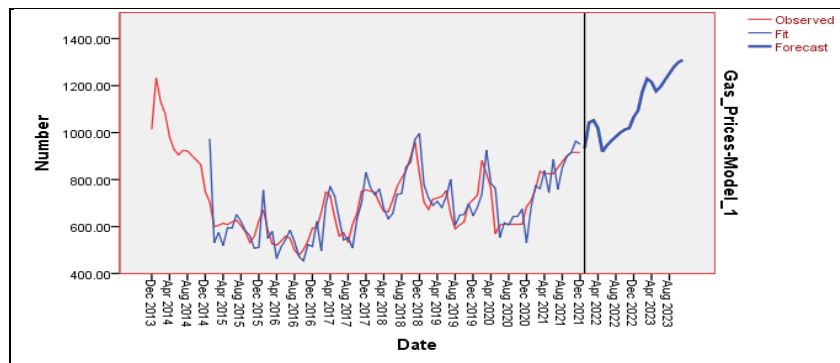


Figure 7: Actual and Estimated of LPG Prices

Table 6: Forecasting for LPG Prices

Year	Forecast	UCL	LCL	Year	Forecast	UCL	LCL
Dec 2021	900.20	995.53	725.50	Dec 2022	1066.72	1674.10	459.34
Jan 2022	933.13	1069.19	797.08	Jan 2023	1094.64	1747.54	441.74
Feb 2022	1042.85	1267.27	818.42	Feb 2023	1177.64	1879.86	475.42
Mar 2022	1052.38	1339.12	765.64	Mar 2023	1229.50	1977.80	481.20
Apr 2022	1018.50	1356.30	680.80	Apr 2023	1215.03	2006.73	423.33
May 2022	918.89	1300.90	536.88	May 2023	1176.55	2009.40	343.70

Jun 2022	945.51	1367.1 6	523.8 7	Jun 2023	1194.7 1	2066.76	322.66
Jul 2022	965.38	1423.2 4	507.5 2	Jul 2023	1223.9 1	2133.47	314.34
Aug 2022	982.81	1474.2 3	491.3 9	Aug 2023	1251.7 3	2197.33	306.14
Sep 2022	999.97	1522.8 0	477.1 5	Sep 2023	1279.4 4	2259.74	299.13
Oct 2022	1012.9 2	1565.3 7	460.4 7	Oct 2023	1299.7 4	2313.56	285.92
Nov 2022	1019.1 4	1599.7 1	438.5 8	Nov 2023	1308.7 5	2355.02	262.48

Conclusion

The results showed that prices would not remain stable throughout the year. The most appropriate SARIMA model for LPG price forecasting of data was found to be SARIMA(0,1,1)(1,1,0)₁₂. From the temporal data, it can be found that forecasted prices would increase to Rs. 1,308.75 in November 2023 from Rs.900.20 in December 2021 in Chennai. The results will be useful to the farming community and other relevant stakeholders. Using time series data from December 2013 to November 2021 on LPG prices, this study provides an evidence on future gas prices in the country, which can be considered for future policy making and formulating strategies for augmenting and sustaining prices in India.

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